



Comparison of performance of exponential, Cox proportional hazards, weibull and frailty survival models for analysis of small sample size data

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Abstract

Background: The Cox Proportional Hazards (PH) model is commonest survival data model used in clinical trials. However, little is known as to whether or not this approach is robust to the size of the study. We compared performance of Cox, Weibull, Exponential and Frailty models in a randomized study that had a small sample size in general.

Methods: Models were fitted to data from a three-arm randomized efficacy trial of Kaposi's sarcoma at Queen Elizabeth Central Hospital (QECH) of children aged <16 years. Patients were randomized to receive Vincristine monotherapy or Etoposide or Vincristine plus Bleomycin. The Cox, Weibull, Exponential and Frailty models were fitted to the data to obtain the adjusted effect of treatment on survival. The performance of models were compared using the Akaike's Information Criteria (AIC). Plots of log-log of survival against log of survival time, and the Schoenfeld's global tests were used to test suitability of the PH assumption.

Results: Ninety-two patients were available for analyses of which 64% were males, the mean age was 8 years (SD=2.8years) and 89% were HIV positive. Exponential model was the best fitting method with AIC=174.1. Children treated with Vincristine monotherapy survived evidently poorly compared with those on Etoposide (HR=5.8, p=0.04).

Conclusion: Exponential models can elicit more valid results than semi-parametric CoxPH model in a clinical trial with small sample size. This emphasizes the fact that when models are fitted to data, it is good practice to assess the goodness of fit and where appropriate alternative models should be fitted.

Keywords: Cox proportional hazards models, survival, parametric models, small sample size, Malawi

Introduction

The Cox proportional hazards (PH) model is the most widely used approach in survival analysis of clinical trials because it requires few assumptions [1]. The model was proposed by Cox in 1972 for analysis of survival data with and without censoring; for identifying differences in survival due to treatment and prognostic factors in clinical trials. The main assumption of the Cox model is that the hazard ratios of any two people are independent of time, known as the PH assumption. However as observed by Ponnuraja and Venkatesan [2], this is not

always appropriate as sometimes the assumptions do not hold. Unfortunately, in practice researchers often do not test models assumptions for fitted models and this practice does not spare the fitting of Cox PH model.

Furthermore, the Cox model is popular because it is considered to be robust to violation of model assumptions; the estimated hazards are always positive and the hazard ratio is often calculated and reported [2,3].

The CoxPH model is parameterized in terms of the hazard model as $h(t)=h_0(t)\times\exp(\beta_1x_{i1}+\dots+\beta_px_{ip})$ (1),

Where $h_0(t)$ is the baseline hazard, $\exp(\beta_1 x_{i1} + \dots + \beta_p x_{ip})$ represents subject-specific component, $x_{i1}, x_{i2}, \dots, x_{ip}$ are the explanatory variables, for i th individual where $i=1, \dots, n$. As can be observed from the expression (1), the Cox PH model is semi-parametric. Alternative models for survival analysis include the Exponential and Weibull, both of which are parametric. Unlike the CoxPH model, the Weibull model allows more flexibility because the associated hazard rate is not constant with respect to time. Also, we use a maximum likelihood process to estimate the unknown parameters. The Weibull model has two parameters, shape parameter (λ); and scale parameter $\frac{1}{\sigma}$. Basing on notation by Hosmer and Lemeshow [2] its survivorship function is

$$S(t|x_i, \beta, \lambda) = \exp\{-t / \exp(\beta_0 + \beta_1 x_{i1} + \dots + \beta_p x_{ip})^\sigma\} \quad (2),$$

while the hazard function has the form

$$h(t, x_i, \beta, \lambda) = \lambda t^{\lambda-1} / \exp\{-(\beta_0 + \beta_1 x_{i1} + \dots + \beta_p x_{ip})^\lambda\} \quad (3),$$

where $\lambda=1/\sigma$, for convenience. The hazard function can be expressed in proportional hazard form as

$$h(t, x_i, \beta, \lambda) = h_0(t) \times \exp(\vartheta x) \quad (4),$$

where the baseline hazard function $h_0(t) = \lambda \gamma t^{\lambda-1}$;

$$\gamma = \exp(-\beta_0/\sigma) = \exp(\vartheta_0) \text{ and } \vartheta_0 = -\beta_0/\sigma.$$

If expressed in the accelerated failure time form the hazard function is given as

$$h(t, x_i, \beta, \lambda) = \lambda \gamma (t \times \exp\{-(\beta_1 x_{i1} + \dots + \beta_p x_{ip})\})^{\lambda-1} \times \exp\{-(\beta_1 x_{i1} + \dots + \beta_p x_{ip})\} \quad (5)$$

A value of $\lambda < 1$ indicates that the failure rate decreases over time, $\lambda = 1$ indicates that the failure rate is constant over time (reduces to an Exponential model) while $\lambda > 1$ indicates that the failure rate increases with time. When the index (shape) parameter equals 1, Weibull reduces to Exponential model. The standard models including the Cox, Exponential or Weibull presented above assume a homogeneous population with independent and identical failure times. To account for heterogeneity of patients in clinical trials caused by unmeasured covariates, extensions of the Cox model called frailty models are introduced. The hazard function of frailty models is given as

$$h(t, z, x) = \vartheta \times h_0(t) \times \exp(\beta^T X) \quad (6),$$

where ϑ is the unobservable random variable. The survival function has the form

$$S(t|\vartheta, x) = \exp(-\vartheta \int_0^t h_0(s) ds \cdot \exp(\beta^T X)) \quad (7),$$

where $S(t|\vartheta, x)$ is the fraction of individuals surviving the time t after follow-up given the vector of observable covariates X and frailty ϑ [2]. In situations where the Cox PH model is deemed inappropriate, it is crucially important to investigate appropriate survival model(s) that can elicit more valid results from the clinical trial. This study emphasizes the importance of assessing model adequacy, where appropriate to consider alternative models, then compare the different models and chose the best fitting model. We compare the performance of the Cox PH, and other PH parametric survival models (Exponential, Weibull and Frailty) on a small sample size data from a Kaposi's sarcoma clinical trial at Queen Elizabeth Central Hospital, Blantyre Malawi. The exponential and the Weibull models were considered because they are the only ones that can be parameterized as either a proportional hazards model or an accelerated failure time (AFT) model. We also evaluated the Frailty model to assess effects of any unmeasured patients' characteristics.

Materials and methods

Study setting and design

The data was from an open label, randomized three-arm trial conducted in the Queen Elizabeth Central Hospital (QECH), Blantyre Malawi, and it compared the efficacy of courses oral Etoposide (reference groups), Vincristine monotherapy, and a combined Vincristine plus Bleomycin therapy. Results for the trial were published [4].

Study population and sample size

The QECH is a 1200 bedded government referral hospital. The children department admits 25,000 children a year and over 90,000 attend the emergency department and outpatients. The children's oncology unit has 25 beds and admits 340 new cases every year. Children with KS are treated as outpatients unless they are systemically unwell. Randomization was computer generated in blocks of 12. Treatment schedules were in sealed, numbered envelopes and opened sequentially at enrolment. Children were enrolled aged <16 years with clinical evidence of Kaposi's sarcoma. As an interim analysis of the trial, this study used a sample of 92 patients, the number collected at the onset of these analyses against 258, estimated during trial design.

Statistical methods

Distributions of baseline characteristics are summarized using numbers with percentages for categorical variables and have been presented by treatment groups. Continuous variables are summarized using medians and interquartile range (IQR). Chi-squared test of independence or K-Wallis test was used to compare distributions between groups for categorical and continuous variables respectively. The Cox, Weibull, Exponential and Frailty survival models were fitted to the data. The survival time was defined as time from randomization to death or censoring date. The outcome

of interest was all-cause death. Patients were censored at their last visit before being lost to follow up for those who dropped out of the study or at the end of the study for those who completed the study still alive. When the sample size is small, Burnham & Anderson (2002) [5] has recommend using second-order Akaike Information Criteria (AICc) over other criteria. However, when all the models in the candidate set have the same number of parameters, then AICc and Akaike Information Criteria (AIC) will give relatively identical estimations. In that situation therefore, AIC can always be used and thus the AIC was used to compare the fit of the survival models in this study. The model with the lowest AIC value is considered to be the best data fitting model. The validity of the PH assumption was assessed using the plots of log-log of survival against log of survival time and the Schoenfeld's global tests were also used. The treatment on which an individual was randomized to was the main exposure of interest. In multivariable analyses, the estimates of the effect of treatment were adjusted for age, gender and HIV status. All analyses were performed using STATA SE version 12(College Station, Texas). Statistical significance was declared at α -level of 0.05. The 95% confidence intervals (CIs) have been provided where appropriate.

Ethical approval

Ethical approval to conduct the randomized controlled clinical trial (main study) was provided by College of Medicine Research and Ethics Committee (COMREC), an independent scientific and ethics committee of the University of Malawi. To uphold confidentiality of the patients, any variables that would be linked to the identity of the participants were removed from the database that was available for this analysis.

Results

Descriptive statistics

There were 92 patients with data available for analysis. Out of the 92 patients, thirty-one were randomised to Etoposide, 31 to Vincristine and 30 to Vincristine plus Bleomycin. There were more males (59) than females (33). However the gender distribution was similar across the treatment groups [4]. Overall, the prevalence of HIV was very high in this study (89%) which was evenly distributed across treatment groups. The mean age was 8 years (SD=2.8years) and was similar across the groups. Across treatment groups, similar distributions were observed in hemoglobin level, absolute neutrophil count, white blood cell count, platelets count, CD4 count and Lansky score, as proportions in oedema, vomiting, tingling, breathless, chest pains, nausea and constipation. Thus randomization was fairly uniform amongst the three treatment arms at least for these known potential confounders.

Main analyses

Model estimation results

In multivariable analyses, the four models yielded different

results regarding effects of treatment on the survival of children. The lowest AIC was obtained from the Exponential model, AIC=174.7 (Table 1) indicating that Exponential model provides best fit to this data compared to the Cox, Weibull and Gamma Frailty models. Overall, the three parametric models had lower values of AIC than the Cox PH model. The value of the AIC for the Cox PH model is substantially higher than the parametric models indicating poor performance.

Table 1. Akaike's Information Criteria (AIC) values obtained from each model.

Model	AIC
Cox	252.4
Exponential	174.7
Weibull	176.7
Shared Gamma Frailty	175.1

Table 1 indicates the lowest AIC of 174.7 obtained from the Exponential model.

In both univariate and multivariate analyses, the four models lead to different statistical inferences in both the crude and the adjusted analyses (Tables 2-5). After adjusting for age, gender and HIV status; children treated with Vincristine monotherapy had a significantly increased hazard of death compared with those on Etoposide (HR 5.8, $p=0.04$; Exponential model in Table 2); (HR 5.7, $p=0.04$; Weibull model in Table 4). The shape parameter in Weibull model did not significantly differ from 1 ($p=0.75$), suggesting insignificant improvement on the Exponential distribution. In contrast, the Cox and Gamma Frailty models showed that the difference in survival of children treated with Vincristine monotherapy and those on Etoposide was not statistically significant at $\alpha=0.05$ (HR 5.0, $p=0.06$; Table 3) and (HR 5.3, $p=0.08$; Table 5) respectively. For the Gamma Frailty model, the random parameter $\vartheta=0.5$, $\log(\vartheta)=-0.7$ and p -value=0.45 suggesting that the random effect is not significant in the model. There was no difference in survival of Etoposide treated children compared with those who received Vincristine combined with Bleomycin in multivariable analyses of all the four models. Age, gender and HIV status did not affect survival of children as depicted from multivariable analyses in all models (Tables 2-5).

Assessment of PH assumption

As can be seen in Table 6, none of covariates in the model as well as the overall Schoenfeld's global test of the full model violates the PH assumption. The plot log-log of survival against log of survival time for study groups indicates roughly parallel graphs for Etoposide and Vincristine monotherapy, thus PH assumption is met in Figure 1. However, there was an interaction between treatment group and time indicating possible violation of PH assumption. The graphs of females and males are roughly parallel as those of HIV positive and negative are, demonstrating no violation of the PH assumption (Figures 2 and 3).

Table 2. Hazard Ratios for death in children with Kaposi's sarcoma from Exponential Model.

Parameter	Univariate model			Multivariable model		
	Hazard Ratio	P-value	95% Estimate CI	Hazard Ratio	P-value	95% Estimate CI
Study drug: Etoposide	Reference					
Vincristine	2.5	0.03*	1.11, 5.39	5.8	0.04*	1.12, 30.21
Vincristine+Bleomycin	1.2	0.70	0.51, 2.73	3.6	0.10	0.78, 16.97
Gender, female	Reference					
Male	0.6	0.08	0.29, 1.08	1.6	0.56	0.32, 8.24
Drug*gender						
Vincristine *male	--	--	--	0.4	0.33	0.05, 2.73
Vinc+Bleomycin*male	--	--	--	0.2	0.08	0.02, 1.25
Age:	1.1	0.22	0.97, 1.15	1.1	0.57	0.99, 1.13
HIV status: Negative	Reference					
Positive	3.5	0.09	0.83, 14.31	2.6	0.23	0.54, 12.83
CD4 count	0.9	0.25	0.99, 1.00	--	--	--

Vinc=Vincristine; bleo=Bleomycin

Table 3 shows that Exponential model suggests a significant increased hazard of death for children treated with Vincristine monotherapy compared with those on Etoposide.

Table 3. Hazard Ratios for death in children with Kaposi's sarcoma from Cox PH Model.

Parameter	Univariate model			Multivariable model		
	Hazard Ratio	P-value	95% Estimate CI	Hazard Ratio	P-value	95% Estimate CI
Study drug: Etoposide	Reference					
Vincristine	2.3	0.04*	1.04, 5.11	5.0	0.06	0.96, 26.19
Vincristine+Bleomycin	1.1	0.86	0.46, 2.51	3.5	0.11	0.76, 16.50
Gender, female	Reference					
Male	0.6	0.20	0.33, 1.26	1.8	0.46	0.36, 9.30
Drug*gender						
Vincristine *male	--	--	--	0.4	0.42	0.06, 3.25
Vinc+Bleomycin*male	--	--	--	0.2	0.08	0.02, 1.25
Age:	1.1	0.54	0.94, 1.12	1.1	0.80	0.92, 1.12
HIV status: negative	Reference					
Positive	3.1	0.13	0.73, 12.72	2.6	0.24	0.53, 12.43
CD4 count	0.9	0.37	0.998, 1.00	--	--	--

Vinc=Vincristine; bleo= Bleomycin

Table 2 shows that the Cox model suggests that survival of children treated with Vincristine and those on Etoposide was essentially not different.

As shown in **Figure 4**, plots of children aged 5-8 years, 9-12 years and 13-16 years do not intersect, proving proportional hazards across these ages. However it is observed that the curve for 1-4 years intersects with 9-12 years and 13-16 years indicating possible violation of PH assumption.

Discussion

The emphasis of these findings is that it is a good practice to assess the goodness of fit of the model and compare with other models. It should be appreciated that more general methods

may not exist and that different methods may perform well under different settings. The methods that have performed well in our setting may prove inadequate in a different setting.

In our model comparisons, we found the Exponential model to have better performance than the other models in the study. Of course we note that the difference in AIC values is small for the Exponential, Shared Gamma Frailty and Weibull models. AIC recommends choosing the model with the lowest AIC value without stating the magnitude of difference. However, for the Cox model, the AIC value is way far from the rest of the models

Table 4. Hazard Ratios for death in children with Kaposi's sarcoma from Weibull Model.

Parameter	Univariate model			Multivariable model		
	Hazard Ratio	P-value	95% Estimate CI	Hazard Ratio	P-value	95% Estimate CI
Study drug: Etoposide	Reference					
Vincristine	2.4	0.03*	1.07, 5.24	5.7	0.04*	1.09, 30.20
Vincristine +Bleomycin	1.2	0.74	1.50, 2.68	3.6	0.10	0.78, 16.92
Gender, female	Reference					
Male	0.6	0.10	0.30, 1.11	1.6	0.55	0.32, 8.28
Drug*gender						
Vincristine *male	--	--	--	0.4	0.34	0.05, 2.80
Vinc+Bleomycin*male	--	--	--	0.2	0.08	0.02, 1.25
Age:	1.1	0.27	0.96, 1.14	1.1	0.58	0.93, 1.13
HIV status: Negative	Reference					
Positive	3.4	0.09	0.81, 14.05	2.6	0.23	0.54, 12.81
CD4 count	0.9	0.11	0.99, 1.00	--	--	--
Shape(σ)	--	--	--	1.0	--	0.75, 1.28
Scale($1/\sigma$)	--	--	--	1.0	--	0.78, 1.33
log(σ)	--	--	--	0.0	0.89	-0.29, 0.25

Vinc=Vincristine; bleo=Bleomycin

Table 4 shows that Weibull model suggests a significant increased hazard of death for children treated with Vincristine monotherapy compared with those on Etoposide. The shape parameter, σ is not significantly different from 1.

Table 5. Hazard Ratios for death in children with Kaposi's sarcoma from Shared Gamma Frailty Model.

Parameter	Univariate model			Multivariable model		
	Hazard Ratio	P-value	95% Estimate CI	Hazard Ratio	P-value	95% Estimate CI
Study drug: Etoposide	Reference					
Vincristine	2.8	0.07	0.91, 7.25	5.3	0.08	0.844, 33.05
Vinc+ Bleomycin	1.1	0.86	0.38, 3.15	4.1	0.11	0.74, 22.69
Gender: female	Reference					
Male	0.6	0.19	0.24, 1.34	1.8	0.53	0.30,10.81
Drug*gender						
Vincristine *male	--	--	--	0.5	0.57	0.05, 5.09
Vinc+Bleomycin*male	--	--	--	0.1	0.08	0.01, 1.30
Age:	1.1	0.69	0.91, 1.16	1.1	0.67	0.91, 1.15
HIV status: Negative	Reference					
Positive	5.0	0.06	0.96, 26.40	3.8	0.15	0.62, 22.74
CD4 count	0.9	0.25	0.99, 1.00	--	--	--
θ	--	--	--	0.5	--	0.08, 3.06
Log(θ)	--	--	--	-0.7	0.45	-2.54, 1.12

Vinc=Vincristine; bleo=Bleomycin

Table 5 shows that Shared Gamma Frailty model suggests that survival of children treated with Vincristine and those on Etoposide was essentially not different.

indicating that its performance was not good in this scenario. Bradburn et al., [6] evaluated the adequacy of some parametric models and the Cox PH model. They found that the generalized gamma model and parametric models achieved both a higher log-likelihood and a lower AIC than the Cox model.

In our study we investigated suitability of the Weibull model by studying the shape parameter. In our findings, Weibull did not improve estimation of the basic Exponential rather was inferior to Exponential model. Nardi et al., [7] also studied Weibull model based on the estimated variation of the

Table 6. Test on proportional-hazards assumption on covariates and full model.

Time:Time	Rho	χ^2	Df	P-value
Vincristine	0.07455	0.19	1	0.6655
vinc+bleo	0.00609	0.00	1	0.9717
Male	0.03202	0.04	1	0.8513
Male*vinc	-0.089558	0.28	1	0.5944
Male*vinc+bleo	-0.07601	0.20	1	0.6531
Age	0.16582	1.22	1	0.2688
HIV positive	-0.30721	3.49	1	0.0619
Global test	--	4.53	7	0.7173

Vinc=Vincristine; bleo=Bleomycin

Table 6 indicates that none of covariates in the model as well as the overall global test of the full model violates the PH assumption.

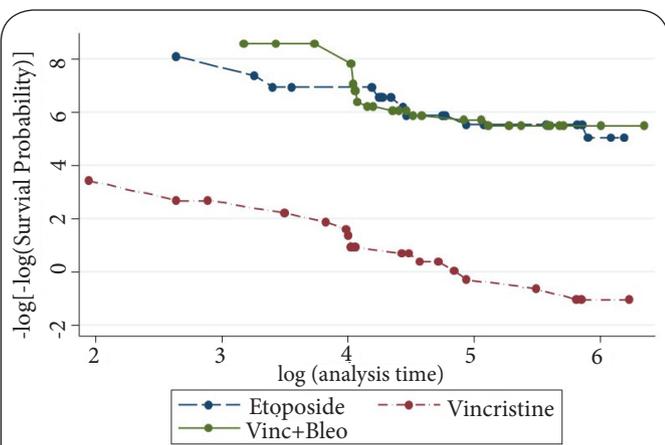


Figure 1. Testing PH assumption on study drug adjusted for age, gender & HIV status.

The plot indicates roughly parallel graphs for Etoposide and Vincristine monotherapy suggesting that PH assumption is met. But graphs of Etoposide and Vincristine plus Bleomycin cross each other showing interaction between treatment group and time indicating possible violation of PH assumption.

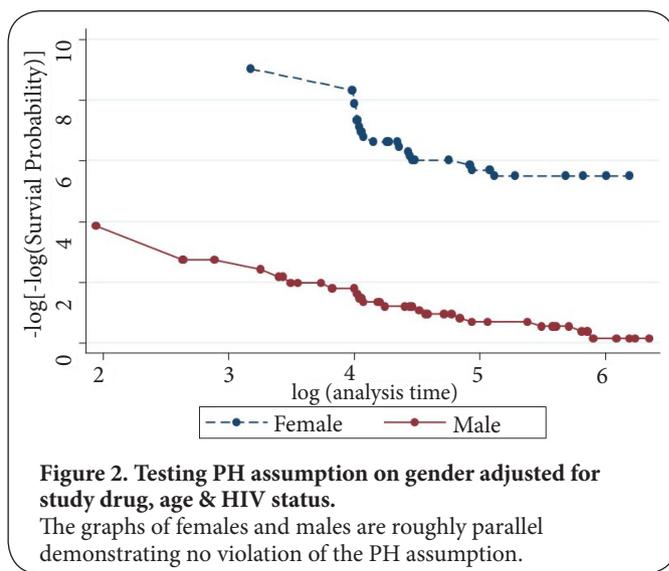


Figure 2. Testing PH assumption on gender adjusted for study drug, age & HIV status.

The graphs of females and males are roughly parallel demonstrating no violation of the PH assumption.

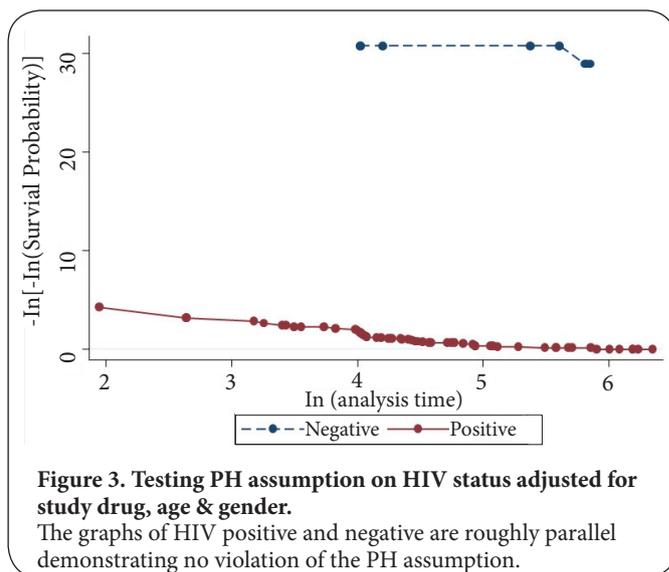


Figure 3. Testing PH assumption on HIV status adjusted for study drug, age & gender.

The graphs of HIV positive and negative are roughly parallel demonstrating no violation of the PH assumption.

parameter rate criteria, and concluded that the Weibull was the superior model.

We explored possibility of the random variation of the survival function using shared gamma frailty model. The standard models including the Cox, Exponential or Weibull assume that a homogeneous population is investigated and their failure times are independently and identically distributed [8]. Henderson and Oman [9] were able to demonstrate that the estimates of regression coefficients may be biased if frailty models are not used where there is frailty effect. Furthermore, Schumacher et al., [10] showed that models that ignore important factors under estimate the relative risk. In our study, hazard functions of the survival models could depend on other unobserved elements and this can result in biased hazard ratios [11,12]. The most common used distributions for

modeling the frailty are gamma or inverse-Gaussian [13-15]. In this paper, we treat frailty as the effect of the unobserved factors on the children's event free survival and conclude that gamma frailty model is not suitable in this study's setting.

A major goal of this paper is to investigate the comparative performance of Cox and parametric survival models by means of real data sets. Results from this study are not reinforced by simulations, which we recommend for further studies. The simulation study would include data simulated under various distributions such as exponential, Weibull, log-normal, log-logistic and with different sample sizes and effect sizes, to evaluate power, type I error rate and precision of the estimated coefficient. Furthermore, in our study we noted some violations of PH assumption. When the PH assumption is violated, we reaffirm that researchers may consider exploring on the

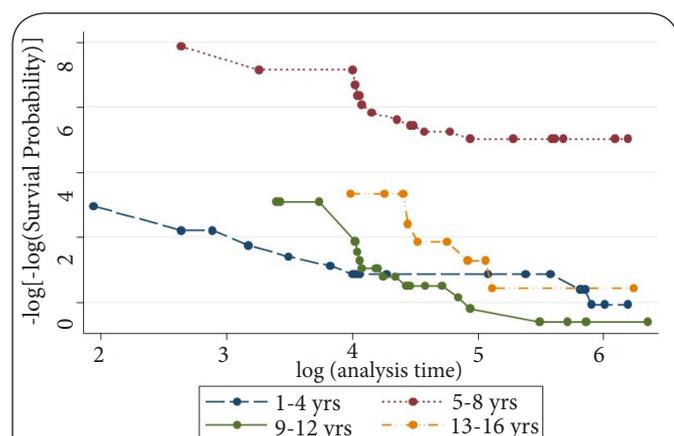


Figure 4. Testing PH assumption on age classes adjusted for study drug, HIV status & gender.

Plots of children aged 5-8 years, 9-12 years and 13-16 years do not intersect, suggesting proportional hazards across these ages, but the curve for 1-4 years crosses with 9-12 years and 13-16 years indicating possible violation of PH assumption.

potential time-dependent models.

Conclusions

More often than not, models are not general in their applications. They are situation specific. In our case we noted that in a clinical trial with small sample size, the Exponential parametric model was the best fitting model compared to the semi-parametric Cox PH, Shared Gamma Frailty and the Weibull models. We recommend and emphasize the very maxim that no matter how routine a method may be used in practice, it is good practice to test the goodness of fit and where applicable it should be compared with alternative models to establish the best fitting model. In some cases the models are fitted by non-statisticians who either do not know how to assess model performance or may not even know that alternative approaches exist. We recommend to such researchers that they should make an effort to verify their analyses with statistician and to help them to explore alternative approaches.

Competing interests

The authors declare that they have no competing interests.

Authors' contributions

Authors' contributions	CS	EM	MM
Research concept and design	✓	✓	✓
Collection and/or assembly of data	✓	✓	✓
Data analysis and interpretation	✓	✓	✓
Writing the article	✓	✓	✓
Critical revision of the article	✓	✓	✓
Final approval of article	✓	✓	✓
Statistical analysis	✓	✓	✓

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References

- Collet D. **Modeling survival data in medical research (2nd Ed.)**. London: Chapman and Hall, CRC. 2003.
- Hosmer D and Lemeshow S. **Applied survival data analysis: Regression modelling of time to event data**. Wiley-Interscience Publication. 1999. | [Pdf](#)
- Ponnuraja C and Venkatesan P. **Survival models for exploring Tuberculosis clinical trial data-an empirical comparison**. *Tuberculosis Research Centre(ICMR)Chennai*. 2010; 3:600. | [Pdf](#)
- Chagaluka G, Stanley C, Banda K, Depani S, Nijram'madzi J, Katangwe T, Israels T, Bailey S, Mukaka M and Molyneux E. **Kaposi's sarcoma in children: an open randomised trial of vincristine, oral etoposide and a combination of vincristine and bleomycin**. *Eur J Cancer*. 2014; 50:1472-81. | [Article](#) | [PubMed](#)
- Burnham KP and Anderson DR. **Model selection and multimodel inference: A practical information-Theoretic approach(2nd ed.)**. Springer-Verlag. 2002. | [Article](#)
- Bradburn MJ, Clark TG, Love SB and Altman DG. **Survival analysis Part III: multivariate data analysis -- choosing a model and assessing its adequacy and fit**. *Br J Cancer*. 2003; 89:605-11. | [Article](#) | [PubMed](#) | [Abstract](#) | [PubMed FullText](#)
- Nardi A and Schemper M. **Comparing Cox and parametric models in clinical studies**. *Stat Med*. 2003; 22:3597-610. | [Article](#) | [PubMed](#)
- Cox DR and Oakes D. **Analysis of survival data**. London: Chapman and Hall. 1984. | [Article](#)
- Henderson R and Oman P. **Effect of frailty on marginal regression**. *J R StatistSoc B*. 1999; 61:367-79. | [Article](#)
- Schumacher M, Olschewski M and Schmoor C. **The impact of heterogeneity on the comparison of survival times**. *Stat Med*. 1987; 6:773-84. | [Article](#) | [PubMed](#)
- Duchateau L and Janssen P. **The Frailty Model**. New York: Inc.: Springer-Verlag. 2008.
- Habil RN. **Frailty models in survival analysis**. *Wittenberg: Halle-Wittenberg*. 2007; 101-108. | [Pdf](#)
- Hougaard P. **Analysis of Multivariate Survival Data**. New York: Springer-Verlag, Inc. 2000; 215-262. | [Article](#)
- O'Quigley J and Stare J. **Proportional hazards models with frailties and random effects**. *Stat Med*. 2002; 21:3219-33. | [Article](#) | [PubMed](#)
- Ghadimi M, Mahmoodi M, Mohammad K, Zeraati H, Rasouli M and Sheikhfathollahi M. **Family history of the cancer on the survival of the patients with gastrointestinal cancer in northern Iran, using frailty models**. *BMC Gastroenterol*. 2011; 11:104. | [Article](#) | [PubMed Abstract](#) | [PubMed FullText](#)

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